## Wage differentials, rent sharing, and worker allocation between firms

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#### More productive firms pay higher wages



9,248,697 total worker-year observations. Data points are averages of all observations allocated each year to a specific percentile. Wage residuals are calculated from an OLS model that controls for year-education-gender dummies and dummies for expected experience. All measures defated by CPI. How has firm productivity contributed to between-firms wage differentials in Finland?

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• Pass-through vs. worker sorting

What drives worker sorting on firm productivity?

- Initial allocation and subsequent reallocation of workers.
- Employment margin responses by firms.

- Pass-through from value-added shocks to wages is small compared to between-firms wage differentials.
- Using firm-switchers for identification, the differentials mostly reflect worker sorting.
- The differentials exist already among labor market entrants.
- Workers more likely to stay, less likely to become unemployed in productive firms.
- Regarding firm level pass-through, firms increase average wages through new hires.

Following discussion by Card et al. (2018).

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- Unobserved employee ability (  $\gamma\uparrow$  )
- Measurement error and fluctuations in value-added ( $\gamma \downarrow$ )
- Endogeneity between wages and productivity (γ ↑↓)

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What can be done?

- Differences for firm-stayers.
- Instrumenting productivity shocks.
- Firm and worker fixed effects (AKM).

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#### Ideal experiments

- Randomized permanent cash windfalls (RCT)
- Firm grant cutoffs (RDD)
- Firm-specific TFP shocks (IV)

- Matched employer-employee data for 1995-2017 (Statistics Finland)
- Sample:  $\sim 500,000$  wage observations per year.
  - Private sector full-time workers in firms with  $\geq 10$  workers.
  - Does not include top managers, owners or selected industries (eg. finance).
  - Only firms with positive value added per worker.
- Productivity:

Gross value added per worker

#### • Wages:

Full hourly compensation incl. overtime & bonuses (& without them)

# OLS rent-sharing elasticity: $ln(w_{it}) = \gamma \ ln\left(\frac{VA_{jt}}{N_{jt}}\right) + X'_{it}\beta + \epsilon_{it}$



Full compensation – Contract-based wage + Overume payments + borluses, basic model controls for years-bn-education-year duritimes and a quadratic of estimated work experience. Occupations are defined at 3-adjut occupation level following the isco5 standard and industries are defined at the 2-digit level following the TOL2008 classification by Statistics Finland. Sample size varies with model depending on data availability. N = 8,339,216 for cross-section, N = 3,376,008 for stayers, N = 3,684,046 for IV model.

Wage differentials, rent sharing, and worker allocation be

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Full compensation = Contract-based wage + Overtime payments + Bonuses. Basic model controls for years-of-education-year dummies and a quadratic of estimated work experience. Occupations are defined at 3-digit occupation level following the Isco5 standard and industries are defined at the 2-digit level following the TOL2008 classification by Statistics Finland. Sample size varies with model depending on data availability. N = 8,339,216 for cross-section, N = 3,376,008 for stayres, N = 3,084,045 for IV model.

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#### Wages of firm-switchers more sensitive to productivity differences



#### Controlling for unobserved sorting using AKM fixed effects

$$\begin{cases} \text{Cross-sectional OLS:} & \ln w_{it} = X'_{it}\beta + \gamma \ln \frac{VA_{jt}}{N_{jt}} + \epsilon_{it} \\ \text{AKM:} & \ln w_{it} = \alpha_i + \psi_{J(i,t)} + X'_{it}\theta + \nu_{it} \end{cases}$$

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$$\Rightarrow \alpha_{i} + \psi_{J(i,t)} + X'_{it}\theta = (\gamma_{\alpha} + \gamma_{\psi} + \gamma_{X\theta}) \ln \frac{VA_{jt}}{N_{jt}} + X'_{it}\beta + (\epsilon_{it} - \nu_{it})$$
  
where  $\gamma_{\alpha} + \gamma_{\psi} + \gamma_{X\theta} \approx \gamma$ 

. . .

#### Rent-sharing elasticity is 0.04 wrt. AKM firm effects



Regression model controls for a quadratic in experience, gender, and level-of-education-year dummies. N = 11,183,347

$$ln(w_{it}) = \gamma \ ln\left(rac{V\!A_{jt}}{N_{jt}}
ight) + X_{it}'eta + \epsilon_{it}$$

- Is  $\gamma$  a constant?
- What determines  $\gamma$ ?
- Omitting leaving workers.

Thank you!

#### Clear productivity effects in entrant wages



#### Firm average wages correlate with firm productivity...



Each percentile holds an equal number of firms, not equal number of workers. N = 9,318,645

#### ... and so do firm average wage residuals



Percentiles 1, 2, 3, 99 and 100 omitted. Wage residuals are calculated from an OLS wage regression for log total hourly wage controlling for year-education group dummies, quadratic in estimated work experience, and gender. Each percentile holds an equal number of furthers N = 9,318,645

#### AKM firm fixed effects correlate with wage residuals



N = 9.137.090

#### Average wages are higher in more productive industries.



- Estimates of rent-sharing elasticity range between 0.05-0.15. Hildreth and Oswald (1997), Arai and Heyman (2009), Carlsson et al. (2016), Card et al. (2018), Lamadon et al. (2019)
- Worker sorting contributes substantially to between-firms wage dispersion. Abowd et al. (1999), Card et al. (2013), Song et al. (2019), Bonhomme et al. (2019), Simola (2020)
- Quasi-experiments find evidence of rent-sharing behaviour. Van Reenen (1996), Saez et al. (2019), Kline et al. (2019), Kroft et al. (2020)
- My contribution: Using a larger sample and further decomposing rent-sharing estimates.

#### Rent sharing in labour market theory

- $\bullet$  Perfect competition  $\Rightarrow$  Law of one wage
- Wage equations in models with imperfect competition:
  - Ex-ante wage posting (Manning, 2011)  $w = \frac{\epsilon}{1+\epsilon}p + \frac{1}{1+\epsilon}b$
  - Ex-post wage bargaining (Manning, 2011)  $w = \alpha p + (1 - \alpha)b$
  - Idiosyncratic workplace preferences (Card et al., 2018)  $w = \frac{\beta}{1+\beta}T_jf_j\mu_j + \frac{1}{1+\beta}b$
- Rationalizing rent sharing: bargaining power, elasticity of labour supply & firm size, asymmetric information & idiosyncratic preferences.

• A simple wage model: 
$$w_i = b_i + \gamma^* \frac{Q_j}{N_i} + \epsilon_{ij}$$

 $w_i$  wage of employee *i* working at firm *j*.

- $VA_i$  value added of firm j
- $Q_j$  quasi-rent of firm  $j: Q_j = VA_j rK_j bN_j$
- $\begin{array}{ll} b_i & \mbox{outside option of the employee.} \\ N_j & \mbox{number of workers in firm } j \\ \gamma^* & \mbox{rent-sharing elasticity.} \end{array}$

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• Empirical counterpart: In 
$$w_{it} = X'_{it}\beta + \gamma$$
 In  $\frac{VA_{jt}}{N_{it}} + \epsilon_{ijt}$ 

• A simple wage model: 
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 $w_i$ wage of employee *i* working at firm *j*. $VA_j$ value added of firm *j* $Q_j$ quasi-rent of firm *j*:  $Q_i = VA_i - rK_i - bN_i$ 

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• Empirical counterpart: In 
$$w_{it} = X'_{it}eta + \gamma$$
 In  $rac{V\!A_{jt}}{N_{it}} + \epsilon_{ijt}$ 

• As  $V\!A_j > Q_j$ ,  $\gamma > \gamma^*$ .

#### Overtime and bonus payments do not fully explain rent-sharing behaviour



#### Industry productivity drives gaps in wage levels, less changes in wages



Regression model controls for a quadratic in experience, gender, and level-of-education-year dummies, N = 11,183,347

#### Estimates of rent-sharing elasticity using AKM fixed effects



Full compensation = Contract-based wage + Overtime payments + Bonuses. Fixed effects are calculated using the AKM model controlling for a cubic in experience, gender, level-of-education-year dummies and 3-digit occupation. Rent-sharing models control for level-of-education-year dummies, quadratic in exprience and 2-digit industries. N = 9,137,090

#### Wages of firm-switchers more sensitive to productivity changes



#### Estimating rent-sharing elasticity

Rent-sharing elasticity $\gamma$ Indep.: log VA per worker (4y ma)		Basic model	Basic + 2-digit industry	Basic + 3-digit occupation	
Α.	Cross-sectional model (levels)				
1)	OLS: Contract-based wage	0.106	0.068	0.061	
		(0.010)	(0.009)	(0.007)	
2)	OLS: Full compensation	0.130	0.085	0.082	
		(0.015)	(0.011)	(0.010)	
В.	Job-stayers model (changes)				
1)	OLS: Contract-based wage	0.006	0.005	0.006	
-		(0.001)	(0.001)	(0.001)	
2)	OLS: Full compensation	0.014	0.012	0.014	
		(0.002)	(0.002)	(0.002)	
3)	IV: Full compensation	0.098	0.073		
,		(0.065)	(0.020)		

Wages in logarithms. Panel A controls for gender, both panels control for a quadratic in experience and education-year dummies. Standard errors clustered at firm level. Instrumental variable: 6-year change in log revenues per worker (t-4 to t+1). First-stage F-statistic for the IV model: xx,xx. Panel A: N = 8,655,006. Panel B: N = 3,919,844

Rent-sharing elasticity $\gamma$ Indep.: log VA per worker (4y ma)	1998-2002	2002-2006	2006-2010	2010-2014	2014-2017
Log full compensation	0.089	0.099	0.040	0.049	0.103
Log person effects	0.055	0.065	0.012	0.040	0.066
Log firm effects	(0.008) 0.033	0.032	0.015)	(0.008) 0.008	(0.007) 0.032
Log covariates	(0.006) 0.002 (0.001)	(0.006) 0.001 (0.000)	(0.007) 0.001 (0.000)	(0.006) 0.000 (0.000)	(0.005) 0.001 (0.000)

Figures in table are regression coefficients for the independent variable when row header is used as the dependent (LHS) variable. Fixed effects are calculated using the AKM model, controlling for age and fixed effects in level of education, 3-digit occupation, and year-dummies. The AKM model and rent-sharing estimates are calculated for each period independently. Rent-sharing estimation model controls for a quadratic in experience, education-year dummies and 2-digit industry-year dummies. Standard errors clustered at firm level. N = 9,137,090.

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